



# Stat 499: Expectations and Sampling methods

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# What is Sampling?

- the selection of a subset (a statistical sample) of individuals from within a statistical population to estimate characteristics of the whole population
- Keywords: **selection, population, estimation**



# Importance Sampling

- Approximate  $E[f]$  by drawing samples from a “proposal distribution”  $q$ , and correcting appropriately by a weighting ratio.
- Suppose dealing with  $p(z)$  is harder, i.e., we **can’t even evaluate  $p(z)$**  but can only do so up to proportionality constant, and only  $\tilde{p}(z)$  can be evaluated. **We can still apply importance sampling by applying the importance weight.**



# Rejection Sampling

- Need to setup a **proposal function**  $q(z)$  and  $M$ , so that  $Mq(z) \geq \tilde{p}(z)$ , for all  $z$ .
- Simulate  $U \sim \text{Unif}(0,1)$  and candidate  $X \sim q$  from the candidate density.
- Use  $U < \tilde{p}(z) / Mq(z)$  to test if reject the candidate  $X$  or not.



# Bayesian inference

- A method of statistical inference in which **Bayes' theorem** is used to update the probability for a hypothesis as more evidence or information becomes available.
- $P(\theta|D) = (P(D|\theta) \times P(\theta)) / P(D)$
- Here, **P( $\theta$ )** is the **prior**, **P(D| $\theta$ )** is the **likelihood** of observing our result given our distribution for  $\theta$ . **P(D)** is the evidence. **P( $\theta$ |D)** is the **posterior belief** of our parameters after observing the evidence i.e the number of heads .
- Use **P( $\theta$ |D)** to estimate the probability of  $\theta$  given the data.



# Maximum likelihood estimation

- A method of **estimating the parameters of a probability distribution by maximizing the likelihood function**, so that under the assumed statistical model the observed data is most probable.
- $\hat{\theta} = \arg \max_{\theta \in \Theta} \hat{L}_n(\theta; \mathbf{y})$  where  $\theta$  is the estimating parameter,  $\mathbf{y}$  is the given data, and  $L_n$  is the likelihood function.
- In practice, it is often convenient to work with **log likelihood**.



# Case study: Zero - Inflated Poisson (ZIP) sampling model

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# Introduction to the problem

- In practice, we often observe **many zeros than we would predict** from a **Poisson model**.
- We will define a **ZIP model** and then find some unknown parameters of the ZIP model using a Bayesian approach.
- We will use the sampling methods (importance sampling and rejection sampling) to answer questions related to the parameters of interest



# Introduction to dataset

- Given the dataset of fish count, we want to **model how many fish are being caught** by fishermen at a state park.
- Some visitors do not fish, but there is no data on whether a person fished or not.
- Some visitors who did fish did not catch any fish so there are excess zeros in the data because of the people that did not fish.

# Head of data

	nofish <int>	livebait <int>	camper <int>	persons <int>	child <int>	xb <dbl>	zg <dbl>	count <int>
1	1	0	0	1	0	-0.8963146	3.0504048	0
2	0	1	1	1	0	-0.5583450	1.7461489	0
3	0	1	0	1	0	-0.4017310	0.2799389	0
4	0	1	1	2	1	-0.9562981	-0.6015257	0
5	0	1	0	1	0	0.4368910	0.5277091	1
6	0	1	1	4	2	1.3944855	-0.7075348	0

6 rows

In this project, we only consider modeling the count variable. Here is the table and summary of the count variable.

0	1	2	3	4	5	6	7	8	9	10	11	13	14	15	16	21	22	29	30	31	32	38	65	149
142	31	20	12	6	10	4	3	2	2	1	1	1	1	2	1	2	1	1	1	1	2	1	1	1

Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
0.000	0.000	0.000	3.296	2.000	149.000



## Model expression

We are going to denote the count variable data as  $y_1, \dots, y_n$ . Assume that  $y_1, \dots, y_n$  are independent and identically distributed given  $0 < \pi < 1$  and  $\lambda > 0$  according to a zero-inflated Poisson sampling model:

$$\begin{aligned} Pr(y_i = 0 | \pi, \lambda) &= \pi + (1 - \pi) \exp(-\lambda) \\ Pr(y_i = k) &= (1 - \pi) \frac{\lambda^k \exp(-\lambda)}{k!} \quad \text{for } k = 1, 2, \dots \end{aligned}$$



## log likelihood function for the parameters

$$l(\pi, \lambda) = r_0 \ln(\pi + (1 - \pi) e^{-\lambda}) + (1 - r_0) \ln(1 - \pi) + \bar{y} \ln(\lambda)$$

$$r_0 = \frac{1}{n} \sum_{i=1}^n \mathbb{I}(y_i = 0)$$

$$\bar{y} = \frac{1}{n} \sum_{i=1}^n y_i$$



## MLE for the likelihood function

Use the optim function in R:

$$\pi_{\text{MLE}} = 0.5677436$$

$$\lambda_{\text{MLE}} = 7.6252130$$

Notice that the  $\lambda$  calculated here is highly influenced by a outlier (149), thus it is biased. We want to **introduce Bayesian inference to eliminate the influence of outliers** by adding prior to the model later.



Estimate the probability that a new fisher catches no fish using the MLE of the parameters:

$$Pr(y_i = 0 | \hat{\pi}, \hat{\lambda}) = \hat{\pi} + (1 - \hat{\pi}) e^{-\hat{\lambda}}$$

$$= 0.5679546$$

Assuming that the new fisher catches at least one fish, estimate the probability that a fisher catches exactly one fish:

$$Pr(y_i = k) = (1 - \hat{\pi}) \frac{\lambda^k e^{-\lambda}}{k!} \quad \text{for } k = 1, 2, \dots$$

$$= 0.003722854$$



## Bayesian analysis

Suppose the prior distribution be:  $\pi \sim \text{Beta}(\alpha, \beta)$ ,  $\lambda \sim \text{Gamma}(\text{shape} = \theta, \text{rate} = \kappa)$ , and assume independence. Assume that  $\alpha = \beta = \kappa = 1$  and  $\theta = 4$ .

Then the log posterior of  $\pi, \lambda$  up to a proportionality constant shows as below:

$$\ln(\text{posterior}(\pi, \lambda | y)) = l(\pi, \lambda) + (\theta - 1) \ln(\lambda) - \kappa\lambda$$

The MAP estimate:

$$\hat{\lambda} = 4.3798186 \quad \hat{\pi} = 0.5625373$$



## Determine the posterior mean of $\pi$ and $\lambda$ using importance sampling

proposal distribution:  $\pi, \lambda \sim \text{multinormal}(\text{mean} = 0.5, 4, \text{sd matrix} = (0.1, 0; 0, 0.5))$

Generate 10000 pair of data from the proposal distribution, calculate the  $q_{\text{tilda}} = \text{dmvnorm}(\text{param})$

Then calculate the  $r_{\text{tilda}} = \text{posterior} / q_{\text{tilda}} = \exp(\log \text{posterior} - \log(q_{\text{tilda}}))$

$$\lambda.\text{mean} = \frac{\sum(\lambda * r_{\text{tilda}})}{\sum r_{\text{tilda}}}$$

$$\pi.\text{mean} = \frac{\sum(\pi * r_{\text{tilda}})}{\sum r_{\text{tilda}}}$$



## Rejection sampling

Use proposal function  $\pi \sim \text{beta}(2,2)$ , and  $\lambda \sim \text{gamma}(8,2)$ . Let  $\pi * \lambda$  be  $g(x)$  and  $\exp(\text{logposterior})$  be  $f(x)$ . Use optim function to find M is around 10000.

Simulate 10000 candidates  $X \sim g$  from the candidate density.

Simulate  $U \sim \text{Unif}(0,1)$ .

$$\text{If } U \leq \frac{f(x)}{Mg(x)}$$

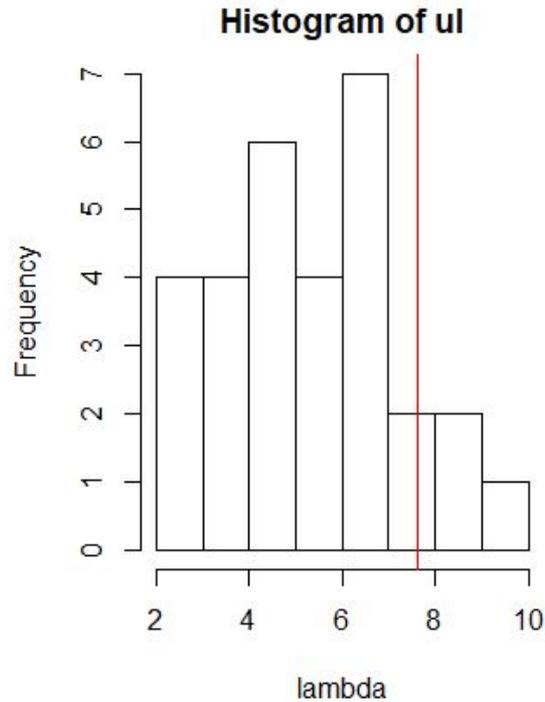
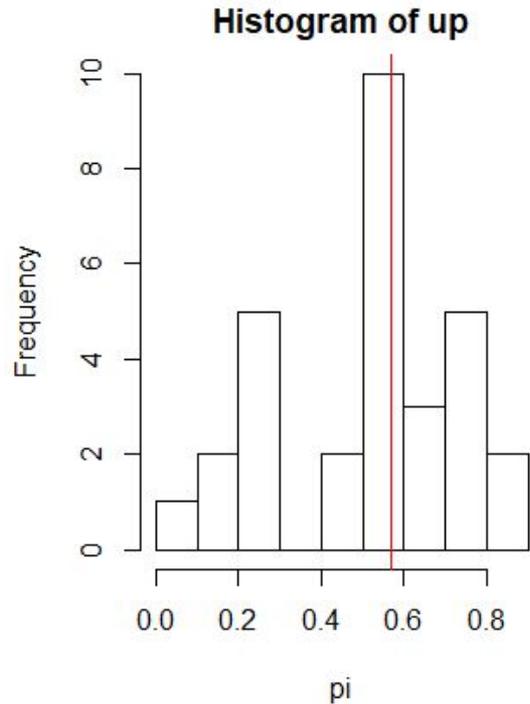
then “accept” the candidate  $X$ . Otherwise, “reject”  $X$ .

# Comparison on statistics

	$\pi$	$\lambda$
MLE	0.5677436	7.6252130
MAP	0.5625373	4.3798186
Importance sampling	0.5242635	4.852539
Rejection sampling	0.5517844	4.875251

We can see that the result of using **importance sampling and rejection sampling does not show significant differences**. They have an overlapped interval of CI for both  $\pi$  and  $\lambda$ . The rejection sampling has a larger standard error because the proposal function can still be improved.

# Visualization of samples using Rejection sampling



The red horizontal line indicates the MLE estimator for the parameters.



## Using the posterior vs naive MLE

	Estimate the probability that a new fisher catches no fish using the MLE of the parameters	Assuming that the new fisher catches at least one fish, estimate the probability that a fisher catches exactly one fish
MLE parameter	0.5679546	0.003722854
Importance sampling	0.5355273	0.07013013
Rejection sampling	0.5227386	0.05840689



# Summary

- In the case study, I used knowledge about Bayes Interface, Maximum likelihood, Importance sampling, and Rejection sampling.
- The result of case study show that these sampling techniques are useful in estimating the expectation of unknown distribution.
- Any questions?